Reflective thinking predicts lower conspiracy beliefs: A meta-analysis

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Abstract

One of the many established predictors of conspiracy beliefs is reflective thinking, but no meta-analysis so far has examined this relationship. In the current meta-analysis of published and unpublished correlational data (145 samples, 181 effect sizes), we found a significant negative association between reflective thinking and conspiracy beliefs with a medium-level effect size (r = -.189). Similar levels of correlations were found across different types of measures (self-report vs. performance-based) and conspiracy beliefs (generic vs. specific). Further, no evidence suggested publication bias in this body of work. Suggestions for future research are discussed. Keywords: analytical, conspiracy, intuitive, meta-analysis, reflective

1 Introduction

Belief in conspiracy theories, alternative explanations for important events that assume the existence of secret plots by powerful and malicious groups (e.g., Goertzel, 1994), have been a common encounter in our social lives for quite some time. For instance, in 2013, it was found that over a third of Americans did not believe that global warming was real (Swift, 2013). Yet, especially with the rise of the COVID-19 pandemic, discussions concerning conspiracy beliefs have become even more prevalent, among psychology researchers as well as laypeople. Growing effort has been put into trying to uncover the psychological mechanisms underlying such beliefs, and rightly so. Belief in conspiracy theories can have far reaching negative consequences for society as a whole as well as for the individual

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believer (Allington et al., 2020; Douglas & Sutton, 2015; Jolley et al., 2019; Jolley et al., 2020), such as reducing willingness to vaccinate against serious diseases and viruses (Bertin et al., 2020; Romer & Jamieson, 2020).

Researchers have revealed several predictors of conspiracy beliefs (van Mulukom et al., 2022). From the personality literature, it is evident that demographic variables such as age (Pizarro et al., 2020; Juanchich et al., 2021), personality traits (Bowes et al., 2020; March & Springer, 2019), and how one reacts to threat and uncertainty (Heiss et al., 2021) influence conspiracy beliefs. From the social domain, although less attended to, research has pointed towards the importance of social context (Alper, 2021; van Prooijen & van Lange, 2014), normative influence (Cookson et al., 2021) and group membership (Cookson et al., 2021a,b; Imhoff & Bruder, 2014). From the cognitive domain, and the focus of this meta-analysis, ample evidence now indicates that reasoning processes and biases such as an individual's tendency to think reflectively can predict their tendency to believe in conspiracy theories (e.g., Alper et al., 2021; Mikušková, 2017; Pytlik et al., 2020; Ståhl & van Prooijen, 2018; Swami et al., 2014). Many studies show that individuals who think reflectively (i.e., analytically, deliberately) are less inclined to believe in conspiracy theories compared to individuals who think intuitively (i.e., experientially), showing that cognitive reflection can act as a buffer against conspiracy beliefs (e.g., Alper et el., 2021; Mikušková, 2017; Pytlik et al., 2020; Ståhl & Van Prooijen, 2018; Swami et al., 2014). However, to date, there has been no attempt to systematically review, meta-analyze or test the probability of publication bias in this body of work. The current meta-analysis aims to address this gap.

1.1 Testing The Link Between Cognitive Reflection And Conspiracy Beliefs

The link between reflective thinking and conspiracy beliefs has been tested in a variety of ways, though mainly via correlational methods. Regarding the assessment of conspiracy beliefs, researchers have applied both generic, such as the 'Generic Conspiracy Beliefs Scale' (GCBS; Brotherton et al., 2013), which capture individuals' generalized tendency toward conspiracy theorizing and more specific measures, such as the 'Belief in Vaccine Conspiracy Theories Scale' (VCBS; Shapiro et al., 2016), which capture belief in a specific set of conspiracy theories. Notably, this body of work has conceptualized reflective thinking in many different ways, typically addressing it as either reflective (Sadeghiyeh et al., 2020), analytic thinking (Swami & Barron, 2021; Stanley et al., 2021), rational thinking (Georgiou et al., 2019) or critical thinking (Lantian et al., 2021). However, they all refer to the same sort of underlying construct, namely, the tendency to engage in reflective and deliberative thought processes as opposed to more automatic and intuitive thought processes (Frederick, 2005; Evans & Stanovich, 2013). For the purpose of uniformity, we refer to this construct as reflective thinking throughout this paper.

In terms of operationalizing reflective thinking, correlational studies have mainly preferred one of two ways – applying either performance-based or self-report measures. The most typical performance-based measures is the Cognitive Reflection Test (CRT; Frederick, 2005). In this test individuals are presented with three different problems. Importantly, there is both an intuitive (low effort) and an analytic (high effort) answer to each problem. For example, for the problem "A bat and a ball cost \$1.10. The bat costs \$1.00 more than the ball. How much does the ball cost?", the intuitive answer occurring to many people would be 10 cents. However, through reflection, it can be calculated that the correct answer is actually 5 cents. Thus, higher scores on the CRT capture an individual's tendency toward reflection, whereas lower scores capture the tendency to think intuitively. Different variations of the CRT have been applied (Thomson & Oppenheimer, 2016; Primi et al., 2016), but importantly, a recurring finding in the literature is that higher scores on the CRT are associated with lower conspiracy beliefs (Adam-Troian et al., 2019; Alper et al., 2021; Sanchez & Dunning, 2021; van Bavel et al., 2022; Wagner-Egger et al., 2018). This has been the case for both generic and specific conspiracy beliefs. For example, Sire et al. (2018) found that participants who reported more correct responses on the CRT (i.e., thought reflectively) reported lower conspiracy mentality and Stanley et al. (2021) as well as Erceg et a.1 (2020) found that higher scores on the CRT were negatively associated with belief in COVID-19 conspiracy theories. Nevertheless, there are also some findings of no significant relationship between scores on the CRT and generic (Ståhl & van Prooijen, 2018) or specific (van Proojen et al., 2022) conspiracy beliefs.

The most commonly used self-report measures consist of variations of the Rational-Experiential Inventory (REI; Epstein et al., 1996) such as the REIm (Norris & Epstein, 2011) or REI-40 (Pacini & Epstein, 1999). Typically, all versions of this scale contain two subscales that capture individuals' preference for reflective (i.e., rational) and intuitive (i.e., experiential) thinking. Importantly, many studies report a significant negative relationship between scores on the reflective thinking subscale and conspiracy beliefs (Swami et al., 2014; Gligorić et al., 2018; 2021), and a positive relationship between scores on the intuitive thinking subscale and conspiracy beliefs (Erceg et al., 2020; Fuhrer & Cova, 2020; Gligorić et al., 2018; 2021; Swami et al., 2014). These findings once again extend to both generic and specific conspiracy beliefs. For example, Georgiou et al. (2019) found that scores on the reflective thinking subscale were significantly and negatively related to generic conspiracy beliefs and, compatibly, Alper et al. (2021) found that scores on the intuitive thinking subscale were positively related to generic conspiracy beliefs. Further, Pytlik et al. (2020) and Tomljenovic et al. (2020) reported findings of a positive relationship between specific sets of conspiracy theories, such as belief in vaccine conspiracy theories, and scores on the intuitive sub scale. Nevertheless, there are also reports of null findings in the literature, for instance Libretto et al. (2020) did not find any evidence for a significant relationship between generic conspiracy beliefs and the rational or intuitive subscale, though their sample size was low compared to other similar studies.

1.2 The Current Study

With growing interest in conspiracy beliefs, the number of studies and, consequently, the number of systematic reviews and meta-analyses in this field have risen in the last few years. Some researchers have systematically reviewed the conspiracy beliefs literature as a whole, attempting to provide a more general state of the art (Douglas et al., 2017; 2019; Goreis & Voracek, 2019; van Prooijen & van Vugt, 2018) or focused on findings regarding a specific set of conspiracy beliefs (van Mulukom et al., 2022). Others, such as Stojanov and Halberstadt (2020) who meta-analyzed the overall effect of lack of control on conspiracy beliefs, have taken more specific findings under the microscope. Yet we know of no effort to systematically review or meta-analyze the currently available evidence. Thus, we conducted a meta-analysis to address the following objectives: (1) identify correlational research that investigates conspiracy beliefs in relation to reflective thinking; (2) identify the different ways in which reflective thinking has been operationalized and measured with respect to conspiracy beliefs; (3) generate a narrative summary of the key findings; (4) assess the risk of publication bias in this work; and overall, (6) summarize findings on the correlational relationship between reflective thinking and conspiracy beliefs, while also providing directions for future research.

2 Method

2.1 Literature search and selection criteria

The literature search was conducted in February and March 2022. As previously recommended (Card, 2015), both discipline-specific databases (Pubmed) and larger, crossdisciplinary databases (Web of Science, Scopus and Google Scholar) were screened. Notably, regarding our searches on Google Scholar, due to the excessive number of results (N = 101000), only the most relevant first 200 articles were screened as previously suggested by Haddaway et al. (2015). The following keywords were used in our search strategy: (conspirac*) AND (rational* OR reflecti* OR intuiti* OR deliberat* OR fast OR slow OR automatic OR system 1 OR system 2 OR thinking style OR bias). Additionally, in order to include unpublished work as well, we screened preprint databases (PsyArXiv and MedRxiv) and sent out an open call for unpublished data via email to all corresponding authors of articles that test the relationship between reflective thinking and conspiracy beliefs. The search log used to record and keep track of all database searches and can be found at OSF (https://osf.io/2jvc3/?view_only=113f7ab290a1415ba9991f65b8b87f1b). To be included in the meta-analysis articles had to contain (1) a quantitative design that statistically tested the correlational relationship between reflective thinking and conspiracy beliefs, (2) a measure and/or manipulation of reflective thinking, (3) a scale measure of conspiracy beliefs, (4) enough information to extract or calculate effect size (e.g., Pearson's r) and (5) English as the language of publication.

The search resulted in a total of 2810 records. Figure 1 shows a PRISMA flowchart of the selection process (Moher et al., 2009). Reviewers (first and second author) firstly screened the titles and abstracts of all collected records. If the article appeared to be relevant, we then assessed the full-text to ensure that it met the inclusion criteria outlined above. This resulted in a total of 267 articles, which then dropped down to 68 after the removal of duplicates. Further, 5 unpublished studies were sent in response to our open call, to which approximately 45% of the contacted authors responded to (regardless of whether they had available data or not). Thus, this left a total of 73 studies and 161 studies to be considered with more scrutiny for a final decision on inclusion in the meta-analysis. In cases where consensus could not be reached between the two reviewers about a paper's eligibility, the third reviewer (senior author) made a final decision. For studies with overlapping data, to avoid using the same data multiple times, simply one study per data set was included in the analysis.



FIGURE 1: PRISMA flowchart of the selection process.

Based on this search strategy, in our final analysis, we included 64 articles (48 published, 16 unpublished), which contained 145 studies (122 published, 23 unpublished) and 181 effect sizes. The data points extracted from these articles were compiled onto a Google spreadsheet (including article characteristics, study characteristics, sample characteristics, analysis characteristics and findings), which can be found on OSF (https://osf.io/bxknp? view_only=113f7ab290a1415ba9991f65b8b87f1b).

2.2 Data Coding and Transformations

To prepare the data for analysis, certain modifications and categorizations were applied to the original dataset. Firstly, the studies were coded based on whether they were published or not (yes vs. no) and, where applicable, the samples' corruption perceptions index (see Supplemental Materials) was added as a new variable to the data set. Secondly, the effect sizes of correlational studies which reported Spearman's rank correlation coefficients were transformed to Pearson's r correlation coefficients using the guidelines recommended by Rupinski and Dunlap (1996) in order to have a common metric. Next, all the available effect sizes were coded as belonging to one of four categories (where "CB" means "conspiracy beliefs" and "RT" means "rational thinking"): (1) Generic CB, Self-Report RT; (2) Generic CB, Performance-Based RT; (3) Specific CB, Self-Report RT; (4) Specific CB, Performance-Based RT. This was done so that, in addition to looking at the overall meta-analytic correlation, we could also meta-analyze the bulk of studies which used selfreport RT measures separately from the ones that used performance-based RT measures, while also separating generic and specific conspiracy belief findings. Finally, two separate datasets were prepared: (1) A weaker estimates dataset and (2) a stronger estimates dataset. Namely, for articles which contained multiple effect sizes in a single study, with more than one effect size that fit into one of our four variable categories mentioned above, the weaker effect size was entered into the weaker estimates dataset (i.e., the main analyses dataset) while the stronger effect size was entered into the stronger estimates dataset. Here we report our meta-analysis findings with respect to the weaker estimates, however, data regarding the stronger estimates can be found in the Supplemental Materials.

2.3 Analysis Strategy

Jamovi's meta-analysis module 'MAJOR' (Hamilton, 2018; version 1.2.0), which is based on the *metafor* package designed for R (Viechtbauer, 2010), was used for our random-effects meta-analysis. The model estimator was restricted maximum-likelihood and model measures were transformed using Fisher's r-to-z transformation. Further, the risk of publication bias, the tendency to omit non-significant results and publish only positive results, was examined in two steps. Firstly, a visual assessment using the funnel plot method (Egger et al. 1997) was conducted, which was followed by a rank correlation test (Begg & Mazumdar, 1994) and a regression test (Egger et al., 1997). Secondly, *p*-curve analysis (Simonsohn et al., 2014) was carried out, using the built-in application available in MAJOR module (Hamilton, 2018). Our study was pre-registered prior to data collection (https://osf.io/ ngxyu/?view_only=fbcc0c15e89f4286bbb7cb469657fc3a).

3 Results

The meta-analysis of 145 independent samples yielded a meta-analytic correlation of r = -.189, SE = .008, z = -22.308, p < .001, 95% CI [-.206, -.172].¹ 94% of all estimates were negative, indicating a clear negative correlation between reflective thinking and conspiracy beliefs (Figure 2). Egger's regression test, value = 2.074, p = .038, but not the rank correlation test, value = .080 p = .155, indicated a funnel plot asymmetry (Figure 3). Note that the asymmetry is opposite to the direction implied by publication bias: the low-powered studies show reduced effects (i.e., lower or positive correlations). The *p*-curve analysis (Figure 4) provided strong support for a true effect with a right-skewed distribution of *p* values. When publication status (yes vs. no) was added to the meta-regression as a categorical moderator, it had no significant association, b = .022, SE = .023, p = .353, 95% CI [-.024, .067], which further suggests that there is no publication bias²

3.1 Self-Report Measures of Reflective Thinking.

When only self-report measures were included, reflective thinking had a significant negative association with generic (k = 28), r = -.173, SE = .035, z = -4.986, p < .001, 95% CI [-.241, -.105] (Figure 5), and specific conspiracy beliefs (k = 23), r = -.146, SE = .025, z = -5.776, p < .001, 95% CI [-.196, -.097] (Figure 6).

3.2 Performance-Based Measures of Reflective Thinking.

When only performance-based measures (e.g., CRT) were included, reflective thinking had a significant negative association with generic (k = 33), r = -.175, SE = .016, z = -10.943 p < .001, 95% CI [-.207, -.144] (Figure 7), and specific conspiracy beliefs (k = 94), r = -.219, SE = .009, z = -24.646, p < .001, 95% CI [-.236, -.201] (Figure 8).

3.3 Summary of Results

The meta-analysis of correlational research suggested a significant negative association between reflective thinking and conspiracy beliefs at r = -.189, which is close to the benchmark of medium effect size, r = .20, in psychological science (Funder & Ozer, 2019).

¹For liberal estimates, the result was similar, r = -.206, SE = .010, z = -21.606, p < .001, 95% CI [-.225, -.188].

²Six Experimental studies were not included in the meta-analysis as there were not enough studies to allow meaningful conclusions. Findings regarding the experimental studies can be found in our supplemental materials: https://osf.io/2cxdg?view_only=113f7ab290a1415ba9991f65b8b87f1b.

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FIGURE 2: Forest plot showing the observed correlations between reflective thinking and conspiracy beliefs. Bars represent 95% confidence intervals. Studies are ranked from the strongest negative to strongest positive correlation. A higher resolution version of the plot is available at https://osf.io/7ejba/?view_only=113f7ab290a1415ba9991f65b8b87f1b.



Fisher's z Transformed Correlation Coefficient FIGURE 3: Funnel plot showing the distribution of associations between reflective thinking and conspiracy beliefs in correlational studies that were published.

Thus, we conclude that reflective thinking negative predicts conspiracy beliefs and the size of this association is sufficient even for practical use in the short run (see Funder & Ozer, 2019). Considering r = -.189, a two-tailed alpha of .05, and a power of 90%, future studies should aim to recruit a minimum of 290 participants to detect a significant correlation between reflective thinking and conspiracy beliefs (Faul et al., 2009).

4 Discussion

This meta-analysis provides empirical and systematic insight to the question: What is the relationship between reflective thinking and conspiracy beliefs? Our analysis of correlational research conducted to date reveals a significant negative association between reflective thinking and conspiracy beliefs at r = -.189, not far from what is considered to be the benchmark of a medium effect size in psychological science (r = .20; Funder & Ozer, 2019). Remarkably, risk of bias assessment via the funnel plot method (Egger et al. 1997) and *p*-curve analysis (Simonsohn et al., 2014) jointly suggests that there is no publication bias in this body of work. Thus, there is sufficient evidence indicating that reflective thinking is a negative predictor of belief in conspiracy theories and that it can be put to practical use to protect individuals from the potentially detrimental outcomes of conspiracy beliefs (Bertin et al., 2020).



FIGURE 4: *p*-curve plot showing the distribution of *p*-values in the published research on the correlation between reflective thinking and conspiracy beliefs. The right-skewed shape suggests that there is no publication bias.

4.1 Self-Report vs. Performance-Based Reflective Thinking Measures

Regarding the measurement of reflective thinking, despite previous research suggesting that the CRT is rather a cognitive ability test rather than a measure of reflective thinking disposition (Otero et al., 2022), we find that both performance based measures (mainly consisting of the CRT) and self-report measures were significant predictors of individuals' belief in conspiracy theories. Namely, we found that both performance-based (r = -.175) and self-report measures (r = -.173) were significantly and negatively associated with generic conspiracy beliefs. Regarding specific conspiracy beliefs, again, we found that both performance-based measures (r = -.146) were significantly and negatively associated with such beliefs.

4.2 Theoretical Background

Why does reflective thinking predict conspiracy beliefs? It is possible that reflectively processing information inhibits individuals' intuitions and biases that support the acceptance

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$\frac{1}{2}$	-0.78 [-0.86, -0.69]
Gkinopoulos et al. (2021) - Study 2 ⊢■⊣	
Sire et al. (2018) ⊢■⊣	-0.52 [-0.61, -0.44]
Georgiou et al. (2018)	-0.32 [-0.42, -0.22]
Georgiou et al. (2019)	-0.32 [-0.42, -0.22]
Fuhrer & Cova (2020) - Study 1	-0.31 [-0.40, -0.21]
Alper et al. (2021) ⊢	-0.29 [-0.35, -0.23]
Stankovic et al. (2021)	-0.29 [-0.40, -0.17]
Fuhrer & Cova (2020) - Study 2 ⊣	-0.24 [-0.37, -0.12]
Stojanov (2015)	-0.21 [-0.37, -0.06]
Lazarevic et al. (2021)	-0.20 [-0.30, -0.11]
Šrol (2022) ⊢■–	-0.19 [-0.29, -0.09]
Gligoric et al. (2018)	-0.15 [-0.33, 0.03]
Gkinopoulos et al. (2021) - Study 1 ⊢∎−	-0.14 [-0.23, -0.05]
Gligoric et al. (2021) ⊢■−	-0.14 [-0.25, -0.04]
Fuhrer & Cova (2020) - Study 3	-0.12 [-0.24, 0.00]
Caroti et al. (2021)	-0.12 [-0.29, 0.05]
Gkinopoulos et al. (2021)	-0.11 [-0.22, -0.00]
Denovan et al. (2020)	-0.10 [-0.20, -0.00]
Atari et al. (2019)	-0.09 [-0.17, -0.01]
Petrovic & Zezelj (2021)	-0.08 [-0.21, 0.04]
Mikušková (2021) ⊢	-0.08 [-0.17, 0.01]
Bensley et al. (2022)	-0.07 [-0.18, 0.04]
Lobato & Zimmerman (2019)	-0.03 [-0.16, 0.10]
Stone et al. (2018)	-0.03 [-0.18, 0.12]
Georgiou et al. (2021)	0.01 [-0.08, 0.09]
Georgiou et al. (2022)	0.01 [-0.08, 0.10]
Fasce & Pico (2019)	0.07 [-0.05, 0.19]
Libretto et al. (2020)	0.09 [-0.05, 0.22]
RE Model	-0.17 [-0.24, -0.10]
-1 -0.8 -0.6 -0.4 -0.2	0 0.2 0.4

FIGURE 5: Forest plot showing the observed correlations between self-report measures reflective thinking and generic conspiracy beliefs. Bars represent 95% confidence intervals. Studies are ranked from the strongest negative to strongest positive correlation.

Georgiou et al. (2021)	■	0.10 [0.02, 0.1
Atari et al. (2019)	↓ ↓ ↓ ↓	0.05 [-0.03, 0.13
Gkinopoulos et al. (2021) - Study 2	-	0.04 [-0.05, 0.13
Georgiou et al. (2022)	├ ── ₽ ──┤	-0.00 [-0.09, 0.09
Gligoric et al. (2021)	├──■ <u></u>	-0.05 [-0.15, 0.05
Cavojová et al. (2020b)	╞──■─┤	-0.05 [-0.16, 0.00
Tomljenovic et al. (2020)	├─■	-0.06 [-0.13, 0.0
Gkinopoulos et al. (2021)	├──■	-0.09 [-0.20, 0.02
Petrovic & Zezelj (2021)	├── ■──┤	-0.16 [-0.29, -0.0
Zezelj & Petrovic (2021)	<u>├</u> ∎]	-0.16 [-0.28, -0.04
Mikušková (2021)	┝╌┲╌┤	-0.18 [-0.27, -0.09
[⊃] ytlik et al. (2020)	┝╌╼╌┤	-0.19 [-0.28, -0.10
Swami & Barron (2021)	┝╌╋╌┤	-0.19 [-0.28, -0.1
Barron et al. (2018)		-0.19 [-0.29, -0.1
Zezelj & Lazic (2019)	►	-0.19 [-0.35, -0.0
Alper et al. (2021)	┝╌╋╌┤	-0.21 [-0.27, -0.1
Swami et al. (2014)	┝╌╋╌┤	-0.21 [-0.28, -0.1
Gligoric et al. (2018)		-0.24 [-0.42, -0.0
Gkinopoulos et al. (2021) - Study 1		-0.24 [-0.33, -0.10
Erceg et al. (2020)	╵	-0.29 [-0.35, -0.2
Feovanovic et al. (2021)	┝──╋──┤	-0.29 [-0.39, -0.1
Georgiou et al. (2019)	⊨	-0.30 [-0.40, -0.1

FIGURE 6: Forest plot showing the observed correlations between self-report measures reflective thinking and specific conspiracy beliefs. Bars represent 95% confidence intervals. Studies are ranked from the strongest negative to strongest positive correlation.

Smelter & Calvillo (2020) Alper et al. (2021) Erceg et al. (2022) - Study 2	├── ₽ -;-1 - ₽ -; -── ₽ -;	-0.09 [-0.23, 0.04 -0.09 [-0.15, -0.03 -0.05 [-0.19, 0.09
Ståhl & van Prooijen (2018) Pennycook et al. (2015)-Study 4 Smelter & Calvillo (2020)		-0.10 [-0.21, 0.07 -0.10 [-0.23, 0.03 -0.09 [-0.23, 0.04
Cavojová et al. (2019) van Prooijen et al. (2022)	┝───₽──┤	-0.11 [-0.29, 0.07 -0.10 [-0.21, 0.07
Marques et al. (2021)		-0.13 [-0.19, -0.07
Fuhrer & Cova (2020) - Study 3 van Prooijen (2017)	↓ . .	-0.14 [-0.26, -0.02 -0.13 [-0.19, -0.07
Marques et al. (2021) ⁄likušková (2021)		-0.14 [-0.21, -0.0] -0.14 [-0.23, -0.0]
3owes & Tasimi (2021) - Study 3 Adam-Troian et al. (2019)	┝╼╋╾┥	-0.16 [-0.22, -0.09 -0.15 [-0.22, -0.09
Vagner-Egger et al. (2018)	⊢ ∎	-0.17 [-0.24, -0.10
Rizeq et al. (2021) Calvillo et al. (2022)	┝── ₽ ──┤ ┊ ├── ₽ ──┤ ┊	-0.18 [-0.29, -0.07 -0.17 [-0.30, -0.04
antian et al. (2021) - Study 2	┝──━──┤	-0.18 [-0.31, -0.0
Calvillo et al. (2021) Fanzer et al. (2021)	┝╌╋╌┤	-0.19 [-0.29, -0.09 -0.19 [-0.26, -0.12
Pisl et al. (2021)	⊢∎⊣	-0.19 [-0.26, -0.13
_ees & Parker (2021) _antian et al. (2021) - Study 1		-0.23 [-0.33, -0.13 -0.20 [-0.42, 0.01
Fuhrer & Cova (2020) - Study 2		-0.24 [-0.37, -0.12
Mikušková (2017)	─── ₽ ──┤ ├── ₽ ──┤	-0.28 [-0.46, -0.10 -0.28 [-0.38, -0.18
Sanchez & Dunning (2021) - Study 1b	├── ■──┤	-0.29 [-0.40, -0.1]
Vagner-Egger et. al. (2018) ⊢	e	-0.31 [-0.47, -0.1
Bowes & Tasimi (2021) - Study 2 Sanchez & Dunning (2021) - Study 1a	· • ·	-0.31 [-0.40, -0.2] -0.31 [-0.43, -0.1

FIGURE 7: Forest plot showing the observed correlations between performance-based measures reflective thinking and generic conspiracy beliefs. Bars represent 95% confidence intervals. Studies are ranked from the strongest negative to strongest positive correlation.



FIGURE 8: Forest plot showing the observed correlations between performance-based measures reflective thinking and specific conspiracy beliefs. Bars represent 95% confidence intervals. Studies are ranked from the strongest negative to strongest positive correlation.

of conspiracy beliefs while also increasing attention to detect the logical fallacies inherent in most conspiracy theories (Brotherton & French, 2013; Swami et al., 2014; van Prooijen & van Vugt, 2018). Further, reflective thinking may equip individuals with the space and skills to override and deliberately evaluate any pre-existing conspiracy beliefs (Swami et al., 2014; Douglas et al., 2019). It could also be the case that conspiracy theories attract an intuitive thinking style simply due to their nature of typically requiring lower cognitive effort and triggering affective feelings (Swami et al., 2014; Swami & Furnham, 2014; van Prooijen & van Vugt, 2018). Finally, it is possible that these findings demonstrate a broader function of reflective thinking, namely, that it fosters skepticism towards any sort of epistemically suspect beliefs (i.e., conspiracy, paranormal and pseudoscience beliefs; Čavojová et al., 2020; Pennycook et al., 2012; Šrol, 2022).

4.3 Future directions

Although the correlational relationship between thinking reflectively and conspiracy beliefs is quite well understood, an evident shortfall of this literature is the lack of high-powered experimental studies. To our knowledge, so far there only exists six studies (Bago et al., 2022; Swami et al., 2014; Spasovski et al., 2022). Thus, the number of experimental studies needs to increase before any informative meta-analytic effect sizes can be produced. Nevertheless, we have added a premature meta-analysis of the currently available studies, with reported Cohen's d effect sizes (N = 4), to our supplemental materials (https://osf. io/2cxdg?view only=113f7ab290a1415ba9991f65b8b87f1b). For future directions, we encourage researchers to engage in more experimental work that tests the effect of reflective thinking on conspiracy beliefs with enough power to detect a significant effect. Such work could include interventional methods such as training individuals to think more reflectively (Isler et al., 2020) or, perhaps, teaching them to reason better (Perkins, 2019). By doing so, one could directly assess whether thinking reflectively does indeed combat against conspiracy beliefs. Further, up to this point only cross-sectional research has been employed and, to our knowledge, there have been no attempts to study this topic via longitudinal methods. Thus, we also encourage researchers to make use of this avenue of research and to study the formation of, or changes in, conspiracy beliefs over time, as a result of increased engagement in reflective thinking.

In regard to sample size, for future reference, we recommend that subsequent studies should aim to recruit at least 290 participants to detect a significant correlation between reflective thinking and conspiracy beliefs based on our finding of r = -.189, a two-tailed alpha of .05, and a power of 90% (Faul et al., 2009). A limitation of the current meta-analyses is that no comparisons could be made regarding study quality characteristics (e.g., pre-registered vs. not pre-registered) due to the current state of the literature being unfit to do so. However, an interesting future direction could be to meta-analyze only pre-registered studies and compare the results with the findings of the current meta-analysis. Relatedly, another limitation of the current study is that no moderators could be assessed,

because most studies in the literature simply reported findings on the relationship between cognitive reflection and conspiracy beliefs. The current meta-analysis supports the idea that this relationship is now well established. Thus, an important step forward would be to start testing the effects of relevant moderator variables, for example corruption (see Supplemental Materials for details), and to further investigate such relations through future meta-analyses, as the number of studies looking at moderators increases. A final potential limitation concerns the p-curve analysis used in the current study. Although popularly applied, the p-curve method has recently faced certain critiques and debates regarding its performance on assessing publication bias, and a new method of assessment, the 'z-curve 2.0', has been proposed (Bartos & Schimmack, 2020). Future studies could utilize this recently developed method to continue to investigate whether it indeed out-performs the more traditionally known p-curve analysis.

4.4 Conclusion

We find that reflective thinking is a significant predictor of conspiracy beliefs, such that individuals who think reflectively are less likely to believe in both generic and specific conspiracy theories. From a theoretical point of view, this paper contributes to the conspiracy beliefs literature by systematically confirming that reflective versus intuitive thinking constitutes one of the underlying psychological mechanisms that support such beliefs. More broadly, from a practical point of view this work suggests that, for policy-makers concerned about the negative outcomes of conspiracy beliefs, encouraging the development of reflective thinking skills can be a good step forward.

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